The Role of Bigram and Unigram Frequencies in Phonotactic Acceptability Judgments

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DGfS
Phonotactic Acceptability

• Categorical:
  – blick
  – *bnick

• Gradient (Halle 1962, Chomsky and Halle 1965):
  – blick > bnick > bzick
Phonotactic Acceptability

• Hypothesis: words that are statistically more probable will be rated higher than words that are less probable.
Findings of the current study

The most obvious means of determining word likelihood actually do not matter when we,

1. control for neighborhood density

2. assess probability based on a model that does not incorporate phonology-specific concepts into the grammar
Modeling gradient judgments

1. Model selection: evaluation metric projected from the lexicon, based on

   - feature-based bigrams (Albright 2009, Heinz & Koirala 2010)
   - bigrams and unigrams combined (Vitevitch et al. 1997)
   - whole word (= neighborhood density) (Bailey and Hahn 2001)
Modeling gradient judgments

2. Parameter tuning:


- maximum likelihood estimation of n-gram models
Modeling gradient judgments

• These models often have built-in phonological components:
  – Features (Albright 2009, Heinz & Koirala 2010)

• Experimental stimuli often have gradient constraint violations:
  – bnick < bzick
Objectives of the current study

1. **Isolate** the role of frequency with a phonotactic acceptability experiment using fully legal CVC nonce words that violate no known phonotactic constraints.

   - no clusters

   - no words beginning with /ŋ/, /ʒ/, /ð/, ending with /h/, etc.
Objectives of the current study

• Previous studies

<table>
<thead>
<tr>
<th>Possible words</th>
<th>Impossible words</th>
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</thead>
<tbody>
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<td></td>
<td>Stimuli</td>
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Objectives of the current study

- Current study

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Objectives of the current study

2. Test whether speakers are sensitive to bigram and/or unigram frequency alone by controlling for neighborhood density.
Bigrams vs. unigrams

• Bigram model: distribution of phonemes in terms of previous sound

Bigrams of ‘blick’ [blɪk]:
   #b, bl, lɪ, ɪk, k#

Bigram probability =
   \( P(b|\#) \times P(l|b) \times P(ɪ|l) \times P(k|ɪ) \times P(\#|k) \)
Bigrams vs. unigrams

- Unigram model: distribution of individual phonemes

  Unigrams of ‘blick’:
  b, l, ɪ, k

  Unigram probability =
  \[ P(b) \times P(l) \times P(ɪ) \times P(k) \]
Training Data

- CELEX English lemma corpus with transcriptions from the CMU Pronouncing Dictionary
- Stress information removed
- Bigram and unigram models trained on resulting corpus of 23,911 words
Experiment

- Forced choice between pairs of nonce words systematically selected to test the individual roles of bigram and unigram frequency.
Stimuli

• All possible CVC words were generated with the English phoneme inventory.

• Real words were removed.

• Bigram and unigram probabilities were determined for each word; zero probability words were removed.

• Each word was given a neighborhood density score (number of real words with a string edit distance of 1) (Jurafsky and Martin 2000)
Stimuli

- Pairs of words selected in five categories - in all pairs, the two words were matched for neighborhood density.
Type 1

• The two words are equivalent* in both bigram and unigram probability.

  – [tʃɔɪb]  vs. [θaʊg]
  BG: 1.15E-08  1.16E-08
  UG: 1.77E-07  1.76E-07

• Prediction: subjects will perform at chance on pairs of this type.
Type 2

- The two words are equivalent in bigram probability, but one word has a higher unigram probability than the other.

  - [nʌs] vs. [bʌʃ]
    - BG: 2.82E-05 vs. 2.82E-05
    - UG: 0.0004 vs. 3.39E-05

- Prediction: subjects will prefer the word with higher unigram probability.
Type 3

• The two words are equivalent in unigram probability, but one word has a higher bigram probability than the other.

  – [sejʃ] vs. [nɪθ]
    BG: 3.95E-05 4.39E-07
    UG: 1.70E-05 1.70E-05

• Prediction: subjects will prefer the word with higher bigram probability.
Type 4

- One word has a *higher* unigram probability but a *lower* bigram probability than the other.

  - [zajp] vs. [dʃeʃ]
    - BG: 2.75E-08 vs. 4.66E-06
    - UG: 5.13E-06 vs. 2.37E-06

- Prediction: ??
Type 5

• One word has a both a higher unigram and a higher bigram probability than the other.

  – [pʌd] vs. [hijð]
    BG: 4.41E-05  2.92E-07
    UG: .0001    3.34E-07

• Prediction: subjects will prefer the word with a higher probability in both models.
Stimuli

• ‘Equivalent probability’ defined as the probability of one word fell into a range of plus/minus 2% of the probability of the other.

• When the words differed in probability, the difference was maximized: one word’s probability fell outside of a range of plus/minus 99.3% of the probability of the other.
Stimuli

- 100 pairs = 200 words total, recorded by a female native speaker of American English
Subjects

- 37 University of Delaware undergraduates, recruited from linguistics courses
- 34 female, 36 right-handed
- ages 18-21
- monolingual speakers of American English
Task

• 100 word pairs were presented aurally in a randomized order.

• Within the pairs, order of presentation for the predicted/not predicted word was counterbalanced.

• “Choose the word you believe is more likely to enter the English language.”
Analysis

• For each stimuli pair type (1-5), calculated the average proportion of trials (out of 20) for which the subjects selected the word predicted by the relevant model.

• Linear mixed-effects models (subject and trial as random effects) were fitted to the data in order to compare the average to chance.
Results

![Bar chart showing average proportion of predicted words selected for different types.]

- Type 1: 0.61
- Type 2: 0.41
- Type 3: 0.46
- Type 4-UG: 0.46
- Type 4-BG: 0.51
- Type 5: 0.51
Results

p values for comparison to chance (α = .05, n = 37)

<table>
<thead>
<tr>
<th>Pair type</th>
<th>p</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>0.02*</td>
</tr>
<tr>
<td>2</td>
<td>0.04*</td>
</tr>
<tr>
<td>3</td>
<td>0.5</td>
</tr>
<tr>
<td>4-UG</td>
<td>0.5</td>
</tr>
<tr>
<td>4-BG</td>
<td>0.9</td>
</tr>
<tr>
<td>5</td>
<td>0.7</td>
</tr>
</tbody>
</table>
Interim summary

• For the pairs in which the models made no prediction, subjects demonstrated a bias for the ‘first word heard’.

• Subjects *disfavored* the words selected by only the unigram model.

• In all other cases, subjects performed at chance.
Model comparison

• Do the previous models correlate better with these results?


Vitevitch & Luce (2004)

• *Positional* unigram probability:

  – CVC

    \[ \text{P(C)} = \text{probability of C appearing in onset position.} \]

• Log frequencies of words with that segment in that position summed and divided by summed log frequencies of all words in the corpus (Merriam-Webster Dictionary)
Vitevitch & Luce (2004)

• Positional bigram probability
  
  – CVC

• Log frequencies of words with those two segments in adjacent positions summed and divided by summed log frequencies of all words in the corpus.
Hayes & Wilson (2008)

- Grammar of phonotactic constraints derived from the training data.
- Nonce word probability calculated based on a weighted sum of its constraint violations.
Heinz & Koirala (2010)

• Distribution of individual feature values (as opposed to phonemes) given the previous value determined from the training data.

• Nonce word acceptability assessed by combining individual feature models into a single model.
Model comparison

- For each stimuli pair, determined the difference in the proportions in which each word was selected:

<table>
<thead>
<tr>
<th>Word 1</th>
<th>% chosen</th>
<th>Word 2</th>
<th>% chosen</th>
<th>Diff.</th>
</tr>
</thead>
<tbody>
<tr>
<td>θuwb</td>
<td>68</td>
<td>θowf</td>
<td>32</td>
<td>36</td>
</tr>
<tr>
<td>nɔs</td>
<td>27</td>
<td>huwb</td>
<td>73</td>
<td>-46</td>
</tr>
</tbody>
</table>
Model comparison

- Likewise, the difference between the scores assigned to the words by each model was calculated:

<table>
<thead>
<tr>
<th>Model</th>
<th>Word 1</th>
<th>Score</th>
<th>Word 2</th>
<th>Score</th>
<th>Diff.</th>
</tr>
</thead>
<tbody>
<tr>
<td>VLUG</td>
<td>nɔs</td>
<td>.1192</td>
<td>huwb</td>
<td>.0875</td>
<td>.0317</td>
</tr>
<tr>
<td>VLBG</td>
<td>nɔs</td>
<td>.0013</td>
<td>huwb</td>
<td>.0022</td>
<td>-.0009</td>
</tr>
<tr>
<td>etc.</td>
<td></td>
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</table>
Model comparison

- Spearman rank correlations were run to determine which model(s) best predicted the experimental results.
## Model comparison

<table>
<thead>
<tr>
<th>Model</th>
<th>ρ</th>
<th>p</th>
</tr>
</thead>
<tbody>
<tr>
<td>Unigram</td>
<td>0.14</td>
<td>0.1</td>
</tr>
<tr>
<td>Bigram</td>
<td>0.07</td>
<td>0.4</td>
</tr>
<tr>
<td>VL-UG</td>
<td>0.06</td>
<td>0.5</td>
</tr>
<tr>
<td>VL-BG</td>
<td>0.02</td>
<td>0.9</td>
</tr>
<tr>
<td>Hayes&amp;Wilson08</td>
<td>n/a</td>
<td></td>
</tr>
<tr>
<td>Heinz&amp;Koirala10</td>
<td>0.19</td>
<td>0.04*</td>
</tr>
</tbody>
</table>
Implications

- If frequency alone can account for gradience (Bybee 1995), then we would expect the subjects to meet the predictions of the bigram and unigram models.
Conclusion and future work

• The current findings suggest that frequency alone is not sufficient to model phonotactic acceptability.

• Further investigation is needed to determine whether this null result is a reflection of the experimental task.
Conclusion and future work

• Additional models can also be compared to better understand what additional information (neighborhood density, syllable structure, featural representations, etc.) should be incorporated.
References

References


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